

MATH 115
LECTURE 17

: Expected Value, Variance / Continuous Random Variables.

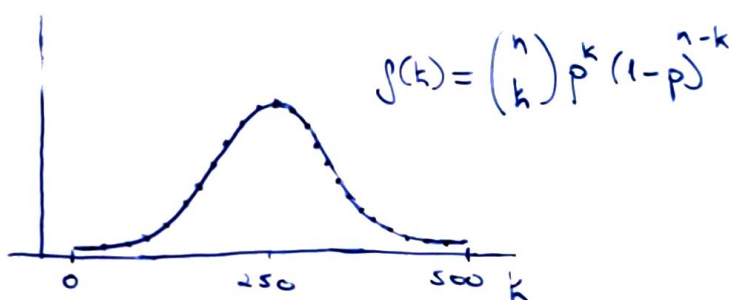
The previous lecture we introduce random variables as a way to study events. Associated to every random variable we have its probability distribution, which summarizes how likely each possible outcome is.

Ex: X = number of heads when tossing a coin 500 times.

↗ so X is a binomial random variable with $n = 500$, $p = \frac{1}{2}$.

Its distribution looks like

$$P(X=k) = f(k)$$



• Moreover, we learn how to compute the expected value of a discrete random variable whose distribution is $f(x)$:

$$E(X) = \sum_k x_k f(x_k).$$

→ In the particular case of a binomial random variable,

$$E(X) = np \quad (\text{expected number of successes in } n \text{ trials, when } p \equiv \text{probability of success}).$$

- Our intuition about the meaning of $E(X)$ is indeed a theorem:

The Law of Large Numbers

Consider an experiment for which we define a random variable X .

(for example: toss a coin 10 times, X = number of heads).

Now, we repeat this experiment many times, and for each time we write down the value of X .

So for example,

1 st experiment	$\rightarrow X_1 = 3$	(3 heads out of 10)
2 nd "	$\rightarrow X_2 = 6$	(6 heads out of 10)
3 rd "	$\rightarrow X_3 = 4$	(4 " " " " ")
:		

If we do it n times, we can define a new random variable, the sample mean of X - $\bar{X}_n$:

$$\bar{X}_n = \frac{X_1 + \dots + X_n}{n}.$$

Then, if we denote $\mu = E(X)$ (expected value of X , also called mean of X), it holds that

$$\lim_{n \rightarrow \infty} P(\mu - \varepsilon \leq \bar{X}_n \leq \mu + \varepsilon) = 1 \quad \text{for all } \varepsilon > 0.$$

(that is, " $\bar{X}_n \rightarrow \mu$ as $n \rightarrow \infty$ ")

- Variance and standard deviation

While $E(X)$ measures the mean value of X , the variance $\text{Var}(X)$ and the standard deviation $\sigma(X)$ measures how much X deviates from its mean $E(X)$. In other words, how likely it is that the values of X are actually near $E(X)$.

Def: Let X be a discrete random variable with distribution $f(x)$ and mean $\mu = E(X)$. Then, we define

$$\begin{cases} \text{Var}(X) = \sum_k (x_k - \mu)^2 f(x_k) = E((X - \mu)^2) \\ \sigma(X) = \sqrt{\text{Var}(X)} \end{cases}$$

Remark: You can show that (check it)

$$\boxed{\text{Var}(X) = E(X^2) - \mu^2}$$

- Binomial random variable: $E(X) = np$

$$\text{Var}(X) = npq, \sigma(X) = \sqrt{npq}.$$

- Example: Find the mean and variance of X .

x	1	3	4	5
$f(x)$	0.4	0.1	0.2	0.3

Sol: $\mu = E(X) = \sum_k x_k f(x_k) = 1(0.4) + 3(0.1) + 4(0.2) + 5(0.3) = 3$

$$\text{Var}(X) = E(X^2) - \mu^2 = \sum_k x_k^2 f(x_k) - \mu^2 =$$

$$= 1^2(0.4) + 3^2(0.1) + 4^2(0.2) + 5^2(0.3) - 3^2 = 12 - 9 = 3.$$

- Exercise: A coin is weighted so that $P(H) = p$ and hence $P(T) = q = 1 - p$. The coin is tossed until a heads appears.

→ Show that the expected number of tosses is $\frac{1}{p}$.

→ Show that the standard deviation in the number of tosses is

Sol:

Sample space: $S = \{H, TH, TTH, \dots, T^k H, \dots, T^\infty\}$

Random variable $X = \text{number of tosses}$, $R(X) = \{1, 2, 3, \dots\}$

Probability distribution: $f(k) = P(X=k) = q^{k-1} p$

Thus,

$$E(X) = \sum_{k=1}^{\infty} k f(k) = \sum_{k=1}^{\infty} k g^{k-1} p = p \underbrace{\sum_{k=1}^{\infty} k g^{k-1}}_{?}.$$

Recall that for any $|g| < 1$ we have that

$$\sum_{k=0}^{\infty} g^k = \frac{1}{1-g}.$$

↓

$$\left\{ \begin{aligned} S_n &= 1 + g + g^2 + \dots + g^n \\ gS_n &= g + g^2 + g^3 + \dots + g^n + g^{n+1} \end{aligned} \right.$$

$$\hookrightarrow S_n(1-g) = 1 - g^{n+1} \Rightarrow S_n = \frac{1-g^{n+1}}{1-g} \Rightarrow$$

$$\Rightarrow \lim_{n \rightarrow \infty} S_n = \frac{1}{1-g} \quad \underline{\underline{=}}$$

Therefore,

$$\frac{1}{pg} \sum_{k=0}^{\infty} g^k = \sum_{k=0}^{\infty} k g^{k-1} = \sum_{k=1}^{\infty} k g^{k-1} = \frac{1}{pg} \left(\frac{1}{1-g} \right) = \frac{1}{(1-g)^2}.$$

and thus we conclude that

$$E(X) = p \frac{1}{(1-g)^2} = \frac{p}{p^2} = \frac{1}{p}.$$

Now,

$$Var(X) = E(X^2) - \mu^2 = \sum_{k=1}^{\infty} k^2 f(k) - \frac{1}{p^2} = \underbrace{\sum_{k=1}^{\infty} k^2 g^{k-1} p}_{?} - \frac{1}{p^2}.$$

We use a similar trick:

$$\frac{d}{dg} \sum_{k=1}^{\infty} k g^k = \sum_{k=1}^{\infty} k^2 g^{k-1} \text{ and we know from the previous}$$

part that

$$\sum_{k=1}^{\infty} k g^{k-1} = \frac{1}{(1-g)^2}, \text{ so } \sum_{k=1}^{\infty} k g^k = \frac{g}{(1-g)^2}, \text{ and thus}$$

$$\sum_{k=1}^{\infty} k^2 g^{k-1} = \frac{d}{dg} \left(\frac{g}{(1-g)^2} \right) = \frac{1}{(1-g)^2} + \frac{2g}{(1-g)^3} = \frac{1-g+2g}{(1-g)^3} = \frac{1+g}{(1-g)^3}.$$

In conclusion,

$$\text{Var}(X) = p \cdot \frac{1+g}{p^3} - \frac{1}{p^2} = \frac{1+g}{p^2} - \frac{1}{p^2} = \frac{g}{p^2},$$

$$\sigma(X) = \frac{\sqrt{g}}{p}.$$

2.1 Continuous random variable.

X is a continuous random variable if its range $\mathcal{R}(X)$ is a continuum of numbers such as $\mathbb{R}$ or an interval.

Say that $\mathcal{R}(X) = [a, b] \subseteq \mathbb{R}$. Then, we associate to X a density function $f(x)$ such that

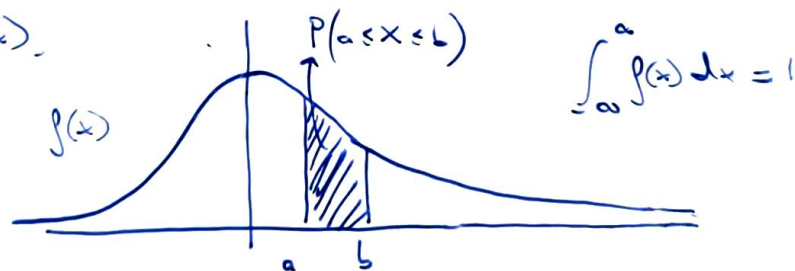
$$P(a \leq X \leq b) = \int_a^b f(x) dx.$$

The function $f(x)$ has to satisfy that

$$1) f(x) \geq 0,$$

$$2) \int_{-\infty}^{\infty} f(x) dx = 1.$$

Remark: A continuous random variable is defined through its density function, and we compute the probabilities as areas under the graph of $f(x)$.



Notice the analogy with discrete random variables:

Here, it is meaningless to talk about the probability of a

particular value of X , since $P(X = x_0) = 0$ for all x_0
 (there are "as possible values for X ").

However, we can consider the probability of being near x_0 :

$$P(x_0 - \frac{1}{2} \Delta x \leq X \leq x_0 + \frac{1}{2} \Delta x) \approx f(x_0) \Delta x.$$

→ In the discrete case:

$$P(a \leq X \leq b) = \sum_k P(X = x_k) = \sum_k f(x_k), \quad \begin{matrix} k = 1, 2, \dots, n \\ x_1 = a, x_n = b. \end{matrix}$$

→ Continuous case:

$$P(a \leq X \leq b) = \int_a^b f(x) dx \approx \sum_k f(x_k) \Delta x_k \approx \sum_k P(x_k - \frac{\Delta x_k}{2} \leq X \leq x_k + \frac{\Delta x_k}{2}).$$

• Def: Cumulative distribution of X

Given X continuous random variable with density function, we define the cumulative distribution of X as

$$F(x) = \int_a^x f(s) ds = P(X \leq x), \quad \text{where } \mathcal{X}(X) = [a, b].$$

(might be $\pm \infty$)

Remark: 1) $P(a \leq X \leq b) = F(b) - F(a)$.

2) $F(x_1) \leq F(x_2)$ for all $x_1 \leq x_2$.

3) $\lim_{x \rightarrow a^-} F(x) = 0, \quad \lim_{x \rightarrow b^+} F(x) = 1$.

• Exercise: Let $f(x) = \frac{1}{x^2}$ on $x > 1$.

a) Show that f is a density function.

b) Find the cumulative distribution of f .

c) Find $P(3 \leq x \leq 4)$.

Sol:

$$a) f(x) \geq 0, \int_1^{\infty} \frac{1}{x^2} dx = \left[-\frac{1}{x} \right]_1^{\infty} = 1.$$

$$b) F(x) = \int_1^x f(s) ds = \int_1^x \frac{1}{s^2} ds = \left[-\frac{1}{s} \right]_1^x = 1 - \frac{1}{x}.$$

$$c) P(3 \leq x \leq 4) = F(4) - F(3) = 1 - \frac{1}{4} - \left(1 - \frac{1}{3} \right) = \frac{1}{3} - \frac{1}{4} = \frac{4-3}{12} = \frac{1}{12}.$$

• Exercise: A point is chosen at random from the unit

square $[0,1] \times [0,1]$. Let X be the sum of its two coordinates.

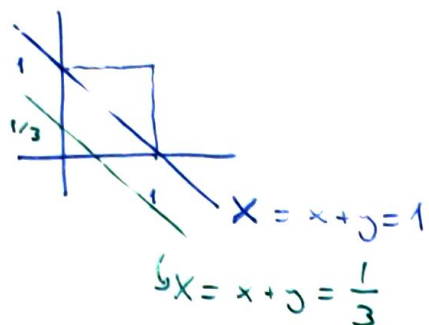
Find the density function and the cumulative distribution of X .

Find $P(\frac{1}{2} < X < 1)$.

Sol: The sample space is $S = \{(x,y) \in \mathbb{R}^2 : 0 \leq x \leq 1, 0 \leq y \leq 1\}$.

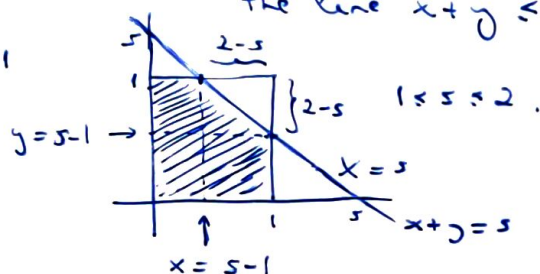
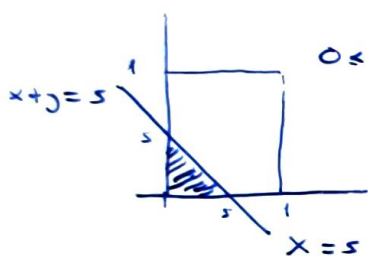
$$\left. \begin{aligned} X : S &\rightarrow [0,2] \\ (x,y) &\mapsto X((x,y)) = x+y \end{aligned} \right\}$$

Notice that X is constant along the lines $x+y=c$: ($X=c$)



Therefore,

$F(s) = P(X \leq s) =$ "area in the square $[0,1] \times [0,1]$ under the line $x+y \leq s$ ".



We can see that

• For $0 \leq s \leq 1$, $F(s) = \frac{s^2}{2}$,

• For $1 \leq s \leq 2$, $F(s) = \frac{1}{2} + \frac{1}{2} - \frac{(2-s)^2}{2} = 1 - \frac{(2-s)^2}{2}$,

that is

$$F(s) = \begin{cases} s^2/2 & 0 \leq s \leq 1 \\ 1 - \frac{(2-s)^2}{2} & 1 \leq s \leq 2 \end{cases}$$

(Check: $F(2) = 1$, $F(0) = 0$).

Then, $f(s) = F'(s)$, so

$$f(s) = \begin{cases} s & 0 \leq s \leq 1, \\ 2-s & 1 \leq s \leq 2, \end{cases}$$

and

$$P\left(\frac{1}{2} < X < 1\right) = F(1) - F\left(\frac{1}{2}\right) = \frac{1}{2} - \frac{1}{8} = \frac{3}{8}.$$

• We can define $E(X)$, $\text{Var}(X)$ and $\sigma(X)$ for continuous random variables analogously to the discrete case:

$$\begin{aligned} E(X) &= \int_a^b x f(x) dx, \quad \text{Var}(X) = E((X-\mu)^2) = E(X^2) - \mu^2 = \\ &= \int_a^b x^2 f(x) dx - \mu^2, \end{aligned}$$

where $\mathcal{X}(X) = [a, b]$ (a, b can be $\pm\infty$),

$$\mu = E(X).$$

• Example. Find μ and σ for X in the previous exercise.

Sol:

$$\mu = E(X) = \int_0^2 s f(s) ds = \int_0^1 s^2 ds + \int_1^2 s(2-s) ds = 1. \quad \left(\begin{array}{l} \text{the result is} \\ \text{intuitive} \end{array} \right).$$

$$\sigma^2 = \text{Var}(X) = \int_0^2 s^2 f(s) ds - 1^2 = \int_0^1 s^3 ds + \int_1^2 s^2(2-s) ds - 1 = \frac{7}{6} - 1 = \frac{1}{6}.$$